

# Projet de Traitement du Signal Segmentation d'images SAR

## Introduction

En analyse d'images, la segmentation est une étape essentielle, préliminaire à des traitements de haut niveau tels que la classification, la détection ou l'extraction d'objets. Elle consiste à décomposer une image en régions homogènes. Les deux principales approches sont l'approche région et l'approche contour. L'approche région cherche à regrouper les pixels présentant des propriétés communes alors que l'approche contour vise à détecter les transitions entre régions. Des détecteurs efficaces ont été développés dans le cadre de l'imagerie optique, mais s'avèrent inadaptés aux images radar de par la présence d'un bruit multiplicatif appelé speckle. L'objectif de ce projet est d'effectuer la segmentation d'une image radar à synthèse d'ouverture (Image RSO ou Image SAR pour Synthetic Aperture Radar) à l'aide d'une méthode originale de détection de ruptures appliquée successivement sur les lignes et colonnes de l'image. La méthode est issue d'une publication intitulée "An Optimal Multiedge Detector for SAR Image Segmentation" publiée en mai 1998 dans la revue *IEEE Transactions on Geoscience and Remote Sensing* [1].

Ce projet s'articule en 4 parties.

### 1) Génération d'une ligne image

Cette partie consiste à générer des lignes d'image radar conformément au modèle proposé par l'article. La méthode de segmentation choisie sera d'abord testée sur ces lignes avant d'être appliquée à des images.

### 2) Analyse spectrale

Dans cette partie, le signal synthétique est étudié à l'aide des outils classiques d'analyse spectrale (corrélogramme, périodogramme). Les courbes obtenues peuvent être comparées aux résultats théoriques énoncés dans la publication.

### 3) Détection des contours

Le détecteur ROEWA est appliqué au signal simulé. Il est basé sur des contrastes locaux de niveau radiométrique moyen.

### 4) Application à des images radar

La segmentation d'une image est réalisée en plusieurs étapes. Les ruptures sont détectées successivement en ligne et en colonne de façon à créer une carte des contours délimitant les régions de radiométrie différente.

## 1<sup>ère</sup> Partie : Génération d'une ligne d'image SAR

### 1) Ligne d'image non bruitée $R(x)$

Une ligne d'image apparaît comme une juxtaposition de segments de réflectivité constante. Elle est correctement modélisée comme un processus constant par morceaux dont les

sauts d'intensité obéissent à un processus de Poisson de paramètre  $\lambda$ . On rappelle que la largeur des segments obéit alors à une loi exponentielle de paramètre  $\lambda$ ,  $\frac{1}{\lambda}$  représentant le nombre moyen de pixels séparant deux sauts d'intensité. (Remarque :  $c$  est le nombre de transitions par intervalle de temps qui suit une loi de Poisson de paramètre  $\lambda$ ). Pour simplifier l'analyse, générer un signal binaire à valeurs dans  $\{1, 2\}$  construit sur un processus de Poisson de paramètre  $\lambda$ . Vérifier que le processus ainsi construit est stationnaire et ergodique à l'ordre un (on déterminera la moyenne arithmétique (spatiale) de chaque réalisation et la moyenne statistique pour chaque pixel  $x_i$ ). On admettra la stationnarité et l'ergodicité à l'ordre 2. Observer et commenter l'influence du paramètre  $\lambda$  sur les signaux générés.

Remarque : la fonction `expnrnd` sous Matlab génère des échantillons suivant la loi exponentielle de paramètre  $1/\lambda$ .

Conformément à l'article [1], générer alors un processus discret à valeurs dans  $\{0, \dots, 255\}$  construit sur un processus de Poisson de paramètre  $\lambda$ .

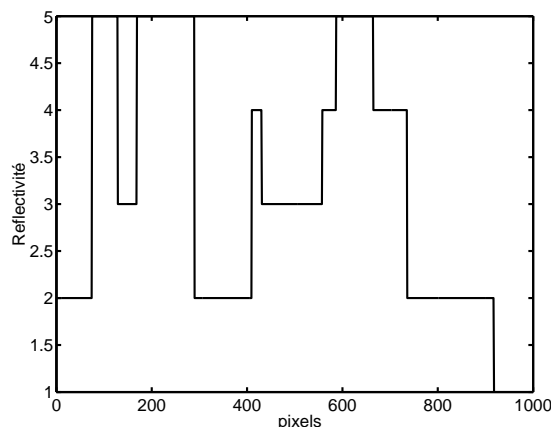


FIG. 1 – Ligne d'image non bruitée

## 2) Bruit Multiplicatif $n(x)$ (Speckle)

Générer une suite de variables aléatoires indépendantes suivant une loi Gamma (fonction `gamrnd`) de moyenne  $\mu_n = 1$  et de variance  $\sigma_n^2 = 1/L$ , où  $L$  correspond au nombre de vues moyennées. Comparer l'histogramme de ces variables aléatoires à la densité de la loi Gamma correspondante (fonction `gampdf`).

Remarque : Matlab considère que la densité de la loi Gamma est :

$$f(x|\alpha, \beta) = \frac{1}{\beta^\alpha \Gamma(\alpha)} x^{\alpha-1} e^{-\frac{x}{\beta}}.$$

## 3) Ligne d'image bruitée $I(x)$

Construire l'intensité d'une ligne d'image Radar à l'aide de  $I(x) = R(x)n(x)$ . Observer l'effet de  $L$  sur l'intensité  $I(x)$  (en considérant par exemple  $L = 1$ ,  $L = 4$  et  $L = 10$ ).

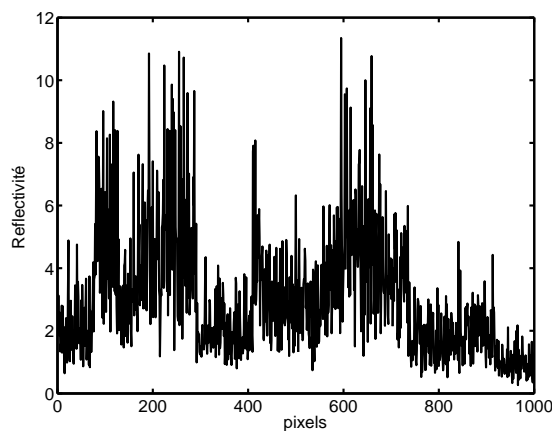


FIG. 2 – Ligne d'image perturbée par un bruit multiplicatif (speckle).

## 2<sup>ème</sup> Partie : Analyse Spectrale d'une ligne d'image SAR

Pour simplifier l'analyse, on commence tout d'abord par étudier le signal binaire à valeurs dans  $\{1, 2\}$  construit sur un processus de Poisson de paramètre  $\lambda$ . On rappelle que la fonction d'autocovariance d'un tel processus s'écrit :

$$C_{RR}(\tau) = E[(R(x) - m_R)(R(x - \tau) - m_R)] = \sigma_r^2 e^{-\lambda|\tau|},$$

où  $m_R = E[R(x)]$  est la moyenne de la ligne d'image. La densité spectrale de puissance est définie comme la transformée de Fourier de la fonction d'autocovariance

$$S_{RR}(f) = \frac{2\lambda\sigma_r^2}{\lambda^2 + 4\pi^2 f^2}.$$

### 1) Périodogramme

Calculer le périodogramme d'une réalisation en utilisant une fenêtre rectangulaire et de Hanning. Comparer le résultat obtenu avec la densité spectrale de puissance associée au processus étudié.

### 2) Périodogramme cumulé

Déterminer la moyenne des périodogrammes de différentes réalisations du signal appelée périodogramme cumulé. Commenter le résultat obtenu.

### 3) Corrélogramme

Déterminer les estimations biaisées et non-biaisées de la fonction d'autocorrélation du processus étudié à l'aide d'une réalisation de ce processus. En déduire une estimation par corrélogramme de la densité spectrale de puissance du processus.

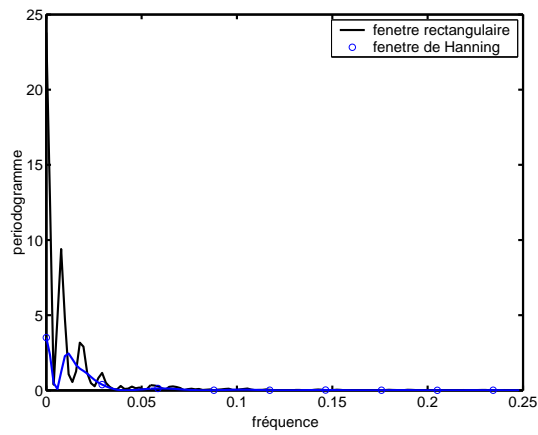


FIG. 3 – Périodogramme

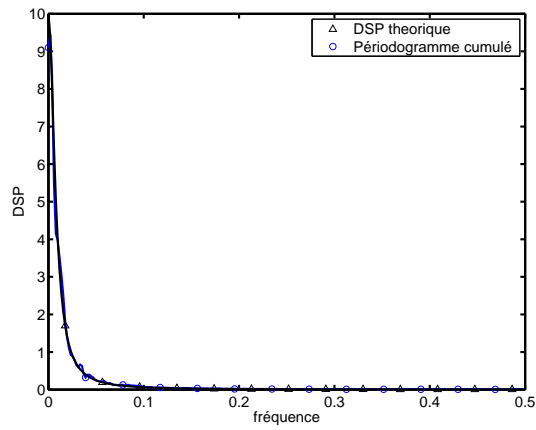


FIG. 4 – Périodogramme cumulé

## Annexes

### Périodogramme/Corrélogramme

La densité spectrale de puissance (DSP) d'un signal à énergie finie est définie par

$$S(f) = TF[K_x(\tau)] = |X(f)|^2$$

où  $X(f)$  est la transformée de Fourier  $x(t)$  et  $K_x(\tau)$  sa fonction d'autocorrélation. Il en découle deux méthodes d'estimation de la DSP appelées **périodogramme** et **corrélogramme**.

– **Périodogramme**

Lorsqu'on estime la transformée de Fourier avec l'algorithme de FFT rapide de Matlab, on montre qu'un estimateur satisfaisant de la DSP du signal  $x(t)$  appelé périodogramme est défini par :

$$\frac{1}{N} |TFD[x(n)]|^2$$

où  $x(n)$  est obtenu par échantillonnage de  $x(t)$ .

– **Corrélogramme**

L'estimation de la DSP par corrélogramme comporte deux étapes :

- 1) Estimation de la fonction d'autocorrélation (xcorr.m) qui produit  $\hat{K}_x(n)$
- 2) Transformée de Fourier discrète de  $\hat{K}_x(n)f(n)$ , où  $f(n)$  est une fenêtre de pondération et  $\hat{K}_x(n)$  est l'estimation biaisée ou non biaisée de la fonction d'autocorrélation.

*Remarque 1* : il est important de noter que lorsque  $\hat{K}_x(n)$  est l'estimateur biaisé de la fonction d'autocorrélation de  $x(t)$ , le corrélogramme coïncide exactement avec le périodogramme comme le montre la figure suivante (obtenue avec une fréquence normalisée  $f_e = 0.1$  :

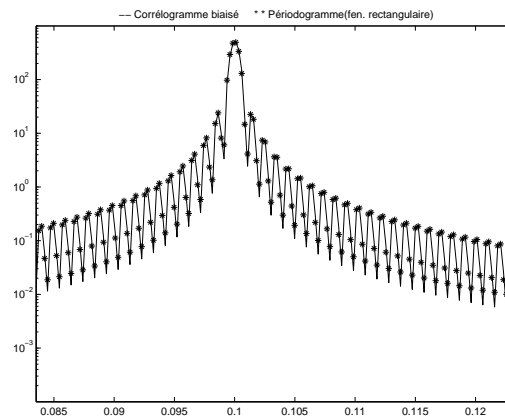


FIG. 5 – Fig. 5. Corrélogramme biaisé et Périodogramme

*Remarque 2 : Implantation numérique*

Si le signal numérique  $x(n)$  possède  $N_s$  points, la fonction xcorr calcule la fonction d'autocorrélation  $\hat{K}_x(n)$  pour  $n = -(N_s - 1), \dots, -1, 0, 1, \dots, (N_s - 1)$  (on a donc  $2N_s - 1$  points). On peut "padder" cette autocorrélation par des zéros afin d'avoir une représentation plus précise de la DSP. L'algorithme de transformée de Fourier discrète de Matlab nécessite une symétrisation de la fonction d'autocorrélation de la façon suivante :

- Points d'autocorrélation  $\hat{K}_x(0), \hat{K}_x(1), \dots, \hat{K}_x(N_s - 1)$
- $N_z$  zéros
- zéro central
- $N_z$  zéros

– Points d'autocorrélation  $\hat{K}_x(N_s - 1), \dots, \hat{K}_x(1)$

Cette procédure de symétrisation est illustrée sur la figure suivante pour  $N_s = 4$  et  $N_z = 4$  :

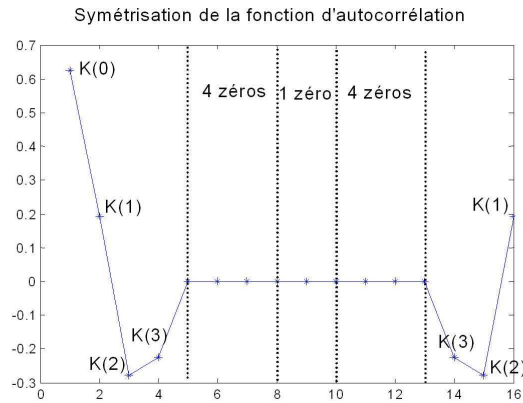


FIG. 6 – Symétrisation de la fonction d'autocorrélation

### 3<sup>ème</sup> Partie : Détection de Ruptures sur une ligne d'image SAR

La détection s'effectue au moyen d'une fenêtre d'analyse glissante. Une forte différence de réflectivité moyenne de part et d'autre d'un pixel permet de repérer un contour. Cette méthode, bien adaptée aux images optiques, est mise en défaut pour l'imagerie radar. La présence d'un bruit multiplicatif augmente en effet le taux de fausses détections dans les régions de forte réflectivité. Pour pallier cette limitation, l'article propose un détecteur basé non plus sur des différences, mais sur des rapports de réflectivité moyenne. En outre, les moyennes arithmétiques sont remplacées par des moyennes pondérées exponentiellement pour traiter les cas de contours multiples (présence de plusieurs contours dans la fenêtre d'analyse). Les plus proches voisins du pixel central sont ainsi favorisés aux dépens de pixels plus éloignés pouvant correspondre à un nouveau contour.

Conformément à l'article [1], on désire réaliser un filtre appelé filtre ISEF de réponse

impulsionnelle

$$f(x) = \frac{\alpha}{2} e^{-\alpha|x|}$$

où  $\alpha$  dépend de  $L$ ,  $\lambda$ , de la moyenne  $\mu_R$  et de l'écart type  $\sigma_R$  de la réflectivité. Pour calculer  $\mu_R$  et  $\sigma_R$  pour un processus à valeurs discrètes  $\{1, \dots, n_r\}$ , on pourra utiliser les formules de sommation :

$$\sum_{k=0}^{n_r} k = \frac{n_r(n_r + 1)}{2}$$

$$\sum_{k=0}^{n_r} k^2 = \frac{n_r(n_r + 1)(2n_r + 1)}{6}.$$

Le résultat du filtrage avec le filtre ISEF est le meilleur estimateur linéaire de la réflectivité au sens de la moyenne quadratique.

On commence tout d'abord par étudier le signal binaire à valeurs dans  $\{1, 2\}$  construit sur un processus de Poisson de paramètre  $\lambda$  bruité par le bruit de scintillement (speckle).

1) Synthétiser le filtre ISEF au moyen d'un filtre RIF. Montrer que le filtrage de  $I(x)$  par ce filtre réalise un débruitage de la ligne de l'image SAR. Commenter ce résultat à l'aide de l'article [1]. Quelques résultats typiques sont représentés sur les figures ci-dessous :

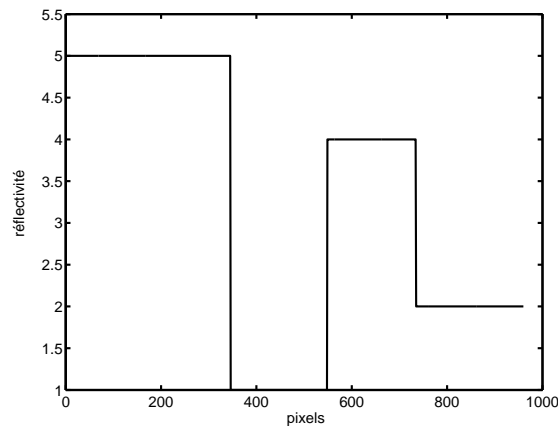


FIG. 7 – ligne non bruitée

2) Conformément à l'article [1], déterminer le rapport des moyennes pondérées exponentiellement noté  $r_{\max}$  (opérateur ROEWA) et montrer qu'il permet d'estimer la position des ruptures.

3) Que se passe-t-il lorsque le signal d'analyse est un processus discret à valeurs dans  $\{0, \dots, 255\}$  construit sur un processus de Poisson de paramètre  $\lambda$  et bruité par du speckle ?

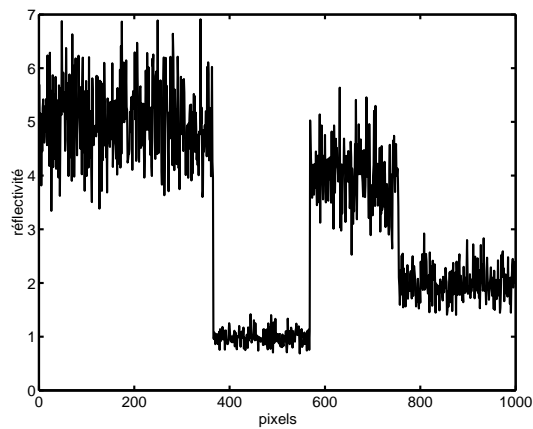


FIG. 8 – ligne bruitée

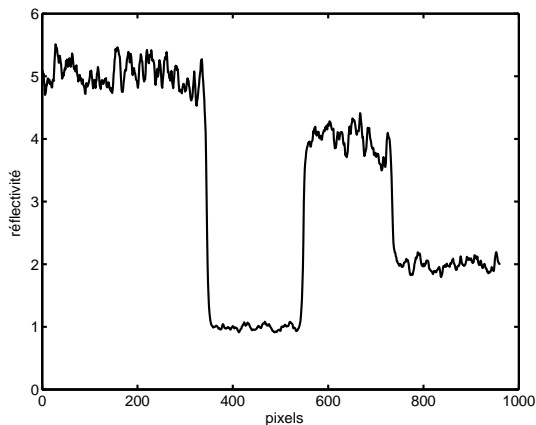


FIG. 9 – ligne débruitée

#### **4<sup>ème</sup> Partie : Détection de Ruptures sur une image TEST**

La détection de contours sur une image est réalisée en deux temps. Le détecteur ROEWA est d'abord appliqué successivement à chaque ligne de l'image, préalablement lissée dans la direction opposée, pour obtenir la carte horizontale des contours  $r_X(x, y)$ . En opérant de façon similaire dans la direction verticale, la carte verticale des contours est construite. Ces deux composantes peuvent alors être combinées pour former la carte des contours en deux dimensions :

$$r_{2D}(x, y) = \sqrt{r_X^2 + r_Y^2}.$$

Retrouver les résultats obtenus sur les images de la figure 5 de l'article [1]. Appli-



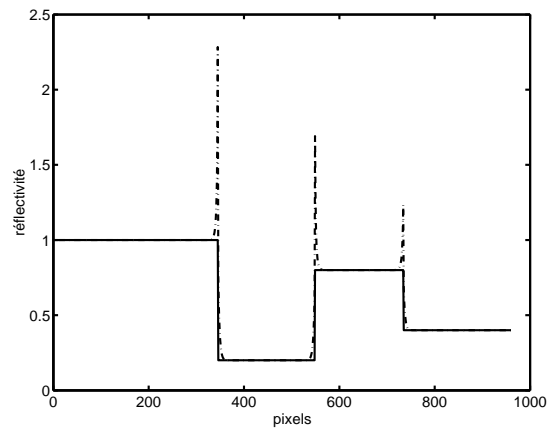


FIG. 10 – Détection de transitions sur une ligne image non bruitée.

quer enfin la méthode à des images non simulées telles que l'image satellitaire de la ville de Bourges (prendre  $L = 4$ ). Pour afficher une image, utiliser "imagesc". Comme on peut le voir, les résultats de segmentation peuvent être décevants dans le cas d'images bruitées. Observer l'effet de la segmentation sur des images synthétiques de Matlab (par exemple, charger l'image eight avec "load indemos eight", transformer les pixels de cette image en réels avec "y=double(eight)") et montrer que les résultats s'améliorent sensiblement.

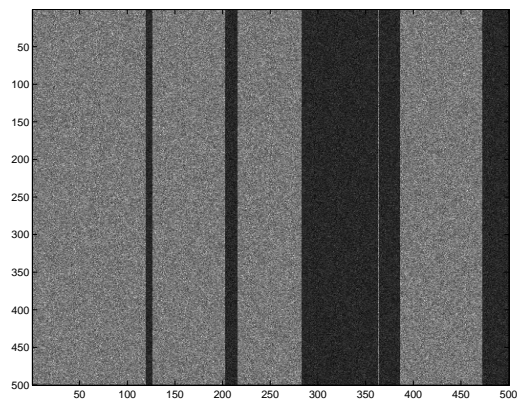


FIG. 11 – Image idéale / lignes verticales

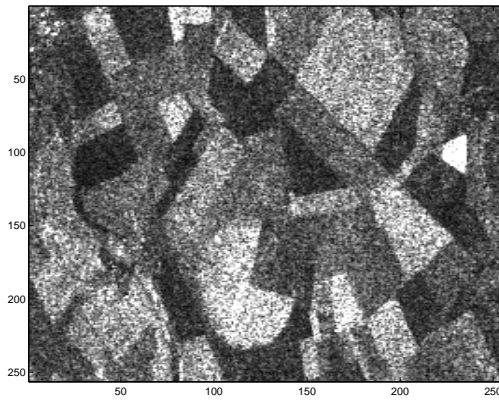


FIG. 12 – Image satellite de Bourges

## Références

- [1] R. Fjørtoft, A. Lopès, P. Marthon and E. Cubero-Castan, “An Optimal Multiedge Detector for SAR Image Segmentation,” *IEEE Trans. Geosci. and Remote Sensing*, vol. 36, n° 3, pp. 793-802, May 1988.

# An Optimal Multiedge Detector for SAR Image Segmentation

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**Abstract**—Edge detection is a fundamental issue in image analysis. Due to the presence of speckle, which can be modeled as a strong, multiplicative noise, edge detection in synthetic aperture radar (SAR) images is extremely difficult, and edge detectors developed for optical images are inefficient. Several robust operators have been developed for the detection of isolated step edges in speckled images. We here propose a new step edge detector for SAR images, which is optimal in the minimum mean square error (MSSE) sense under a stochastic multiedge model. It computes a normalized ratio of exponentially weighted averages (ROEWA) on opposite sides of the central pixel. This is done in the horizontal and vertical direction, and the magnitude of the two components yields an edge strength map. Thresholding of the edge strength map by a modified version of the watershed algorithm and region merging to eliminate false edges complete an efficient segmentation scheme. Experimental results obtained from simulated SAR images as well as ERS-1 data are presented.

**Index Terms**—Edge detection, multiedge model, region merging, segmentation, speckle, synthetic aperture radar (SAR), watershed algorithm.

## I. INTRODUCTION

SEGMENTATION is the decomposition of an image in regions, i.e., spatially connected, nonoverlapping sets of pixels sharing a certain property. A region may for example be characterized by constant reflectivity or texture. Region-based segmentation schemes, such as histogram thresholding and split-and-merge algorithms, try to define regions directly by their content, whereas edge-based methods try to identify the transitions between different regions.

In images with no texture, an edge can be defined as an abrupt change in reflectivity. In the case of optical images,

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an edge is usually defined as a local maximum of the gradient magnitude in the gradient direction, or equivalently, as a zero-crossing of the second derivative in the direction of the gradient. Smoothing is necessary prior to derivation, as differential operators are sensitive to noise. The smoothing and differentiation operations are merged and implemented by two-dimensional (2-D) filters. Gradient-based edge detection basically consists in calculating the *difference* of the local radiometric means on opposite sides of the central pixel. This is done for every pixel position in the vertical and horizontal direction, and the magnitude of the components is computed. Finally, local maxima of the gradient magnitude image are extracted.

Owing to the multiplicative nature of speckle, edge detectors based on the difference of average pixel values detect more false edges in areas of high reflectivity than in areas of low reflectivity in synthetic aperture radar (SAR) images [1]. Certainly, other measures than the difference can be used to identify abrupt transitions. Several edge detectors with constant false alarm rates (CFAR's) have been developed specifically for SAR images, e.g., based on a ratio of averages [1], [2] or a likelihood ratio [3], [4]. However, these operators use the arithmetic mean for the estimation of local mean values, which is optimal only in the monoedge case. Segmentation schemes based on region growing [5], [6], histogram thresholding [7], and simulated annealing [8] have also been proposed for SAR images.

In this article we concentrate on the spatial aspect of edge detection, based on a multiedge model. We incorporate the specific properties of SAR images and develop a linear minimum mean square error (MMSE) filter for the estimation of local mean values. In this way we obtain a new edge detector with improved noise suppression and edge detection properties. Section II explains the principle of monoedge detection in SAR imagery. In section III we develop an optimal multiedge detector and propose a thresholding method which extracts closed, skeleton boundaries. The use of region merging to eliminate false edges is also described. Experimental results obtained from simulated SAR images and ERS-1 data are presented in section IV. We discuss theoretical aspects and experimental results in section V, and end with some concluding remarks in section VI.

## II. MONOEDGE DETECTION IN SAR IMAGES

Speckle is a deterministic effect common to all imaging systems relying on coherent illumination. It is due to the constructive and destructive interference of the responses of the different elementary scatterers of a resolution cell. In the measured intensity image  $I$ , speckle is well modeled as a multiplicative random noise  $n$ , which is independent

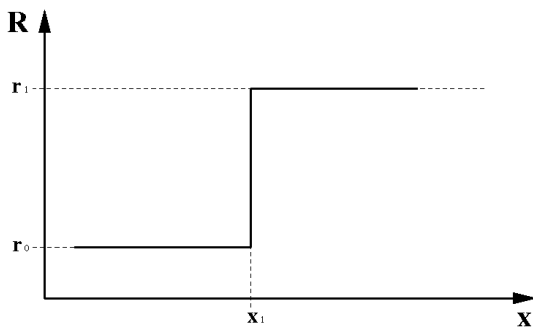


Fig. 1. One-dimensional monoedge model.

of the radar reflectivity  $R$  [9]

$$I(x) = R(x) \cdot n(x). \quad (1)$$

The transfer function of the SAR system is designed to vary as little as possible over the bandwidth of interest. It is known to have negligible influence on the spectrum of the ideal image, but to limit the bandwidth of the noise spectrum. This effect is here incorporated in the term  $n$ . Fully developed speckle is Gamma distributed with mean value  $\mu_n = 1$  and variance  $\sigma_n^2 = 1/L$ , where  $L$  is the equivalent number of independent looks (ENIL) of the image [9].

Several CFAR edge detectors have been developed for SAR images based on the monoedge model, which supposes that only one step edge is present in the analyzing window (Fig. 1). For example, Touzi *et al.* showed that edge detectors based on the Ratio Of Averages (ROA) have CFAR, because the standard deviation  $\sigma_I$  is proportional to the mean intensity  $\mu_I$  [1]. The ratio is normalized to lie between zero and one

$$r_{\min} = \min \left\{ \frac{\hat{\mu}_1}{\hat{\mu}_2}, \frac{\hat{\mu}_2}{\hat{\mu}_1} \right\} \quad (2)$$

where  $\hat{\mu}_1$  and  $\hat{\mu}_2$  are the arithmetic mean intensities of the two halves of a window of fixed size. The normalized ratio is calculated in four (or more) directions, by splitting the analyzing window along the horizontal, vertical and diagonal axes. The minimum of the four values thus obtained is finally compared to an edge detection threshold, which is set according to the accepted probability of false alarm (PFA), i.e., the probability of detecting an edge in a zone of constant reflectivity.

The principle of the likelihood ratio (LR) detector is to estimate the ratio of the probability that the analyzing window covers two regions separated by a given axis to the probability that the entire window belongs to one single region. Transforming the LR for edge detection in SAR images into a log-likelihood difference yields [4]

$$\ell_{\text{edge}} = \nu (-N_1 \log \hat{\mu}_1 - N_2 \log \hat{\mu}_2 + N_0 \log \hat{\mu}_0) \quad (3)$$

where  $\nu$  is the order parameter of the Gamma distribution of the SAR image,  $N_1$ ,  $N_2$ ,  $\hat{\mu}_1$ , and  $\hat{\mu}_2$  are the number of pixels and the arithmetic mean values of the two half windows, and  $N_0$  and  $\hat{\mu}_0$  are the corresponding parameters for the entire window. Oliver *et al.* recently showed that the ROA operator coincides with the LR operator if only the averages are estimated on equally sized halves of the sliding window [4].

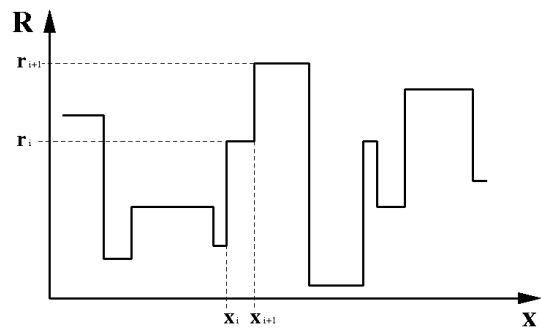


Fig. 2. One-dimensional multiedge model.

The unbiased maximum likelihood (ML) estimator of the mean value of a Gamma distributed *stationary* process, is the arithmetic mean [4]. The ROA and LR operators both use this estimator. It is optimal under the monoedge model, i.e., as long as the width of each half window does not exceed the minimum distance between significant edges. In SAR images the signal to noise ratio is very low, typically 0 dB for single-look images. To sufficiently reduce the influence of the speckle, an important number of pixels must be averaged in each half window. So there is a conflict between strong speckle reduction and high spatial resolution, and the chosen window size constitutes a compromise between these two requirements. This illustrates the limitations of the monoedge model.

### III. MULTIEDGE DETECTION IN SAR IMAGES

For most scene types, the large windows that we use to detect edges in SAR images are likely to contain several edges simultaneously. In fact, we need to estimate the *local mean values*  $\{\hat{\mu}_{r_i}\}$  of a signal which undergoes abrupt transitions with random intervals. The monoedge hypothesis is generally not verified, and the arithmetic mean is no longer optimal. Estimators with nonuniform weighting should therefore be considered. The filter coefficients decide the weighting of the pixels as a function of the distance to the central pixel. For our application, they should optimize the tradeoff between noise suppression and spatial resolution, based on *a priori* knowledge of image and noise statistics.

#### A. Multiedge model

We restrict ourselves to a separable image model. In the horizontal as well as in the vertical direction we suppose that the reflectivity image (ideal image)  $R$  is a stationary random process composed by piecewise constant segments of reflectivity  $\{r_i\}$ , with mean value  $\mu_r$  and standard deviation  $\sigma_r$ . The localization of the reflectivity jumps  $\{x_i\}$  follows a Poisson distribution with parameter  $\lambda$  corresponding to the mean jump frequency, i.e., the probability of  $k$  jumps in the interval  $\Delta x$  is given by

$$p_k(\Delta x) = \frac{(\lambda \Delta x)^k}{k!} e^{-\lambda \Delta x}.$$

The reflectivities  $\{r_i\}$  and the jump localizations  $\{x_i\}$  are supposed to be independent. Hence  $\mu_R = \mu_r$  and  $\sigma_R^2 = \sigma_r^2$ . Fig. 2 illustrates the multiedge model in the one-dimensional (1-D) case. Although it is idealized, this model

is a good approximation for important scene types, such as agricultural fields.

It can easily be shown that the autocovariance function of the reflectivity is [10]:

$$C_{RR}(\Delta x) = \sigma_r^2 e^{-\lambda|\Delta x|}.$$

The ideal image is thus a separable first-order Markov process with parameter  $\lambda$ . The power spectral density, which we here define as the Fourier transform of the autocovariance function, is then

$$S_{RR}(\omega) = \frac{2\lambda\sigma_r^2}{\lambda^2 + \omega^2}. \quad (4)$$

### B. Linear MMSE filter

Let us now develop the linear MSSE filter for the estimation of the local mean under the stochastic multiedge model and the multiplicative noise model. It should not be confused with an adaptive speckle filter [11], which restores the reflectivity of a pixel based on the local statistics. The MMSE filter will be split along the vertical and horizontal axes, and the weighted means estimated in the different half windows will be used for edge detection. To facilitate the implementation, we suppose the filter to have separable impulse response  $f_{2-D}(x, y) = f(x)f(y)$  and first consider the one-dimensional case. The best unbiased linear estimator of the reflectivity is of the form [12]

$$\hat{R}(x) = \mu_R + f(x) * (I(x) - \mu_I). \quad (5)$$

Minimizing the mean square error  $E[|R(x) - \hat{R}(x)|^2]$  yields the transfer function [12]

$$F(\omega) = \frac{\mu_n S_{RR}(\omega)}{S_{RR}(\omega) * S_{nn}(\omega) + \mu_R^2 S_{nn}(\omega) + \mu_n^2 S_{RR}(\omega)}. \quad (6)$$

The autocovariance function of the speckle decreases very rapidly [9]. As an approximation,  $n$  will be considered as white noise here

$$C_{nn}(\Delta x) = \sigma_n^2 \delta(\Delta x)$$

$$S_{nn}(\omega) = \sigma_n^2 = 1/L$$

By substituting the power spectral densities and mean values into (6) and taking the inverse Fourier transform we obtain the optimal impulse response

$$f(x) = C e^{-\alpha|x|}$$

where

$$\alpha^2 = \frac{2L\lambda}{1 + (\mu_R/\sigma_R)^2} + \lambda^2 \quad (7)$$

and  $C$  is a normalizing constant. From the multiplicative noise model (1) we have  $\mu_R = \mu_I$  and

$$\sigma_R^2 = \frac{L\sigma_I^2 - \mu_I^2}{L + 1}$$

which can be estimated from the speckled image. The average region width  $1/\lambda$  can be evaluated visually, or we can

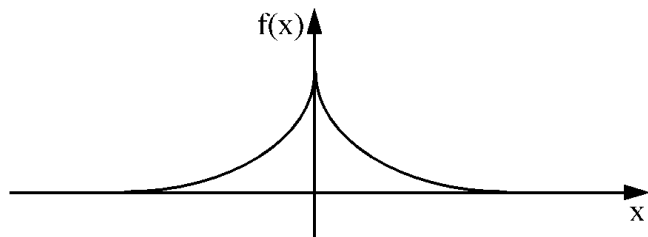


Fig. 3. Impulse response of the infinite symmetric exponential filter (ISEF).

estimate  $\lambda$  from the spectrum of a speckle-reduced image (4) obtained by adaptive filtering [11].

We normalize  $f$  with respect to the mean value, i.e.,  $C = \alpha/2$ , to obtain an unbiased estimator. With this normalization  $f(x) * \mu_I = \mu_I$  and (5) simplifies to

$$\hat{R}(x) = f(x) * I(x).$$

We can thus apply the filter directly to the measured intensity image  $I$ .

The impulse response of  $f(x)$  is shown in Fig. 3. As we see, the filter is of infinite extent, which for the two-dimensional filter  $f_{2-D}(x, y) = f(x)f(y)$  means that the analyzing window centered on the pixel to be filtered covers the entire image. The weight of the surrounding pixels decreases exponentially with distance. The further a pixel is from the center, the more likely it is to belong to another region, and the less influence it has on the estimated local mean. We note that  $f_{2-D}$  is not strictly isotropic.

The filter  $f$  is known as the infinite symmetric exponential filter (ISEF). The ISEF is the basis of the edge detector of Shen and Castan [13], which computes the difference of the exponentially weighted means in each half window. This is an optimal multiedge detector for images degraded by additive white noise. It is claimed to have better edge localization precision than other edge detectors proposed for optical images [13]. We have now shown that the same type of smoothing filter is optimal in the case of multiplicative noise. However, as explained in section I, edge detectors based on the difference of averages are not suited for SAR images.

We also note the analogy between the ISEF and the Frost speckle filter [11]. Frost *et al.* assumed the *local* variations, within stationary regions of the image, to be a first order Markov process and developed an adaptive restoration filter where local statistics control the slope of the exponential weighting function. We use the first order Markov process as a *global* image model for the optimization of a nonadaptive filter.

In the discrete case,  $f$  can be implemented very efficiently by a pair of recursive filters [13], [14]. We define two discrete filters,  $f_1(n)$  and  $f_2(n)$ , realizing the normalized causal and anti-causal part of  $f(n)$ , respectively

$$f_1(n) = a \cdot b^n u(n), \quad (8)$$

$$f_2(n) = a \cdot b^{-n} u(-n), \quad (9)$$

where  $0 < b = e^{-\alpha} < 1$ ,  $a = 1 - b$ , and  $u(n)$  is the discrete Heaviside function. The smoothing function can now be rewritten as

$$f(n) = c \cdot b^{|n|} \equiv \frac{1}{1+b} f_1(n) + \frac{b}{1+b} f_2(n-1)$$

where  $c = (1 - b)(1 + b)$ .

By taking the z-transform of (8) and (9) we obtain

$$\begin{aligned} F_1(z) &= \frac{a}{1 - bz^{-1}} \\ F_2(z) &= \frac{a}{1 - bz} \end{aligned}$$

In terms of the spatial index  $n$ , convolution with  $f_1(n)$  and  $f_2(n)$  corresponds to the following simple recursions

$$\begin{aligned} s_1(n) &= a e_1(n) \\ &+ b s_1(n - 1) \quad n = 1, \dots, N \end{aligned} \quad (10)$$

$$\begin{aligned} s_2(n) &= a e_2(n) \\ &+ b s_2(n + 1) \quad n = N, \dots, 1. \end{aligned} \quad (11)$$

Here  $e_1(n)$  and  $e_2(n)$  are the inputs, and  $s_1(n)$  and  $s_2(n)$  are the outputs of  $f_1$  and  $f_2$ , respectively. To minimize the number of multiplications, we may rewrite (10) and (11) as

$$\begin{aligned} s_1(n) &= a (e_1(n) - s_1(n - 1)) \\ &+ s_1(n - 1) \quad n = 1, \dots, N \\ s_2(n) &= a (e_2(n) - s_2(n + 1)) \\ &+ s_2(n + 1) \quad n = N, \dots, 1. \end{aligned}$$

The computational cost for  $f_1$  and for  $f_2$  is thus one multiplication per pixel. Due to the normalizing factors,  $f$  necessitates four multiplications per pixel.

### C. The ROEWA operator

Based on the linear MMSE filters described above, we propose a new ratio-based edge detector: the *ratio of exponentially weighted averages* (ROEWA) operator. The exponentially weighted averages  $\hat{\mu}_1$  and  $\hat{\mu}_2$  are normalized to be unbiased, and we show in the Appendix that their variance is proportional to the variance of the raw image. The standard deviation remains proportional to the mean value, so the ROEWA operator has CFAR [1]. As opposed to Touzi *et al.*, (2), we normalize the ratio to be superior to one

$$r_{\max} = \max \left\{ \frac{\hat{\mu}_1}{\hat{\mu}_2}, \frac{\hat{\mu}_2}{\hat{\mu}_1} \right\}. \quad (12)$$

The two approaches are of course equivalent. Our choice is motivated by the particular algorithm that we use in the edge extraction step.

To compute the horizontal edge strength component, the image  $I(x, y)$  is first smoothed column by column using the one-dimensional smoothing filter  $f$ . Next, the causal and anti-causal filters  $f_1$  and  $f_2$  are employed line by line on the result of the smoothing operation to obtain  $\hat{\mu}_1(x)$  and  $\hat{\mu}_2(x)$

$$\begin{aligned} \hat{\mu}_{X1}(x, y) &= f_1(x) * (f(y) * I(x, y)) \\ \hat{\mu}_{X2}(x, y) &= f_2(x) * (f(y) * I(x, y)). \end{aligned}$$

Here  $*$  denotes convolution in the horizontal direction and  $\star$  denotes convolution in the vertical direction. The normalized ratio  $r_{X \max}(x, y)$  is found by substituting  $\hat{\mu}_{X1}(x - 1, y)$  and  $\hat{\mu}_{X2}(x + 1, y)$  into (12). The vertical edge strength

component is obtained in the same manner, except that the directions are interchanged

$$\begin{aligned} \hat{\mu}_{Y1}(x, y) &= f_1(y) * (f(x) * I(x, y)) \\ \hat{\mu}_{Y2}(x, y) &= f_2(y) * (f(x) * I(x, y)). \end{aligned}$$

Finally, with analogy to gradient based edge detectors for optical images, we take the magnitude of the two components

$$|r_{2-D \max}(x, y)| = \sqrt{r_{X \max}^2(x, y) + r_{Y \max}^2(x, y)}.$$

In the edge strength map thus obtained, a high pixel value indicates the presence of an edge. For each pixel this implies a total of 14 multiplications, an average of 3 divisions, and 1 square root operation.

### D. Edge extraction

By thresholding the edge strength map we obtain pixels which, with a certain PFA, belong to edges. If the threshold is set too high, we miss important edges, and if it is set too low, we detect a lot of false edges. Plain thresholding will in general produce several pixel wide, isolated edge segments. The edges can be thinned to unity width e.g., using morphological closing [1]. The problem of forming closed boundaries from spatially separated edge segments is quite complicated. If the edges are not closed, they do not define a segmentation of the image.

The *watershed algorithm* [15] is a simple and efficient edge detection method which gives closed, skeleton boundaries. The edge strength map is considered as a surface and the algorithm detects local maxima by immersion simulation. In its original form, the watershed algorithm retains all of the local maxima of the edge strength map, which separate different basins. It unfortunately tends to produce massively over-segmented images. We have chosen to introduce an edge detection threshold in the algorithm [16]. Only edge strength magnitudes over the chosen threshold are considered. Local maxima with lower magnitudes are supposed to be due to noise. With this modification, the algorithm detects, thins and closes significant edges in one operation. The modified watershed algorithm is illustrated in Fig. 4.

We do not have any analytical expression for the distribution of the exponentially weighted means. When the slope of the exponential function is moderate, however, we may suppose a Gaussian distribution according to the central limit theorem. The variance of the distribution as a function of the variance of the raw image, the speckle correlation and the filter parameter  $b$  is given in the Appendix. The relation between detection threshold and PFA can be established theoretically for the ROEWA operator based on the Gaussian hypothesis. In fact, as the Gamma distribution fits a Gaussian distribution very closely when the ENIL is a few tenths or higher, the PFA computed for the ROA operator [1] can also be used for the ROEWA for typical values of  $b$ . The ENIL of the exponentially weighted mean is equal to the ENIL of the raw image multiplied by the equivalent number of independent pixels in the half window, which is given in the Appendix. The PFA applies to the vertical or horizontal edge strength component, but only as an approximation to their magnitude. Moreover,

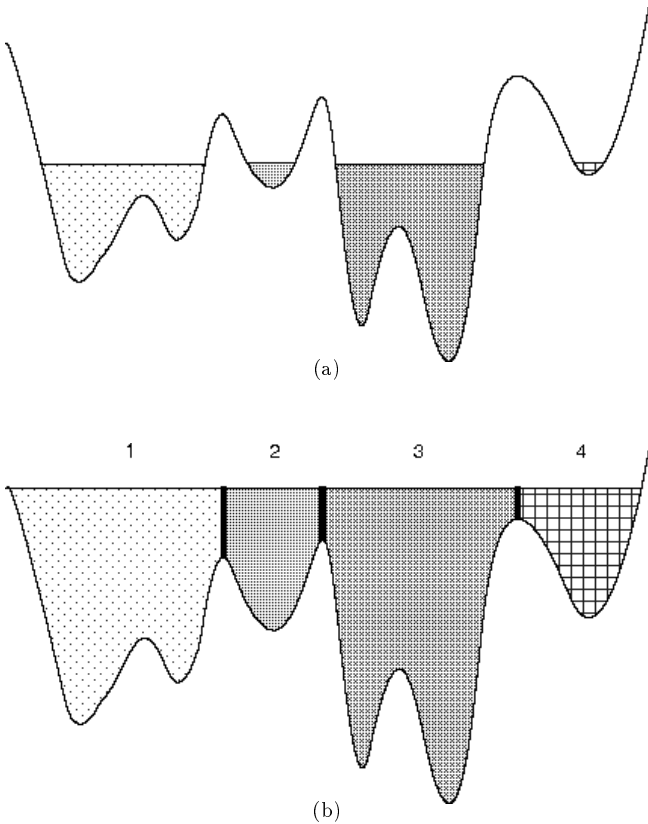


Fig. 4. (a) Initial state of the modified watershed algorithm shown on a cross-section of an edge strength map. (b) Skeleton boundaries detected after complete immersion.

watershed thresholding reduces the PFA compared to plain thresholding, as also false edges are thinned to unity width. The effect of this nonlinear operation is difficult to quantify. With our approach, the theoretical PFA for a given threshold can therefore only serve as a rough indication.

A particularity of watershed thresholding is that the whole edge is eroded if the edge strength magnitude of one single edge pixel is below the detection threshold. The threshold must consequently be set relatively low for the algorithm to form meaningful boundaries, but then we are bound to detect numerous false edges as well.

### E. Post-processing

Spurious edges can be eliminated by merging adjacent regions whose reflectivities are not significantly different. Several merging criteria have been proposed, including the Student's  $t$ -test [6] and the unequal variance Student's  $t$ -test [14]. The LR of Oliver *et al.* [4] can also be used to decide whether or not two regions should be merged, and again constitutes an optimal criterion. In fact,  $\ell_{\text{merge}} = -\ell_{\text{edge}}$  (3)

$$\ell_{\text{merge}} = \nu (N_1 \log \hat{\mu}_1 + N_2 \log \hat{\mu}_2 - N_0 \log \hat{\mu}_0). \quad (13)$$

Thus  $\ell_{\text{merge}} \leq 0$ , and a value close to zero suggests that the two regions together form a Gamma-homogeneous region. It should be noted, however, that we in many applications seek a thematic segmentation, so that weak textures within the regions can be accepted. In practice, negative thresholds are used. The more irregularities we accept

within the regions, the further the threshold can be from zero. Again, the threshold can be related to the PFA [4].

Geometrical considerations, such as region size [14] and edge regularity [6], may also be taken into account in the merging process, based on *a priori* knowledge about the size and shape of the regions. The order in which the regions are merged has a strong influence on the final result. Finding the globally optimal merging order requires much time-consuming sorting. The *iterative pairwise mutually best merge criterion* [17] is a locally optimal approach which is much quicker. First all regions are compared with their neighbors in terms of the merging criterion, and the results are stored in a dynamic array. The array is then traversed sequentially, and a region  $\mathcal{A}$  is merged with an adjacent region  $\mathcal{B}$  if and only if  $\mathcal{B}$  is the closest neighbor of  $\mathcal{A}$  according to the merging criterion, and if  $\mathcal{A}$  is also the closest neighbor of  $\mathcal{B}$ . When two regions are merged, the local statistics of the resulting region must be updated and the comparison with all its neighbors must be redone before continuing. The array is traversed repeatedly until no adjacent regions satisfy the merging criterion.

## IV. EXPERIMENTAL RESULTS

The novelty of our detector is that it relies on weighted means rather than on the arithmetic means used by other CFAR detectors. To study the influence of the nonuniform weighting, we compare the ROEWA operator with the ROA operator. For both detectors the normalized ratio  $r_{max}$  is computed vertically and horizontally, and the magnitude of the two components constitutes the edge strength map. We use the modified watershed algorithm for thresholding, because it directly yields skeleton boundaries, localized on local maxima of the edge strength map. This property facilitates the subsequent tests.

A quantitative comparison of edge detectors can only be effectuated on simulated images, as we need to know the exact position of the edges in advance. Let us first consider a “cartoon image” composed of vertical bands of increasing width, from 2 to 18 pixels. The ratio between the reflectivities of the bright and the dark lines is 12dB. This reference image was multiplied with a simulated single-look speckle image. The correlation coefficients of the speckle is  $\rho(1) = 0.42$ ,  $\rho(2) = 0.03$ , and  $\rho(m) = 0$ ,  $m > 2$ , in azimuth as well as in range. The ideal image and its single-look speckled counterpart are shown in amplitude in Fig. 5 (a) and (b), respectively. Edge strength maps were calculated on the speckled image with both operators. Single-look images are extremely noisy, so strong smoothing is necessary. The ROEWA operator with  $b = 0.9$  produced a very regular edge strength map giving rise to few false edges. To obtain the same reduction of speckle variance with a half window for both operators, and thus the same false alarm rate for a given detection threshold, the window size for the ROA operator was set to  $39 \times 39$  (see the Appendix). A threshold of 1.85 gave the best compromise between the detection of real edges and the suppression of false ones.

The resulting segmentations are shown in Fig. 5 (c) and (d). The ROEWA operator gives a systematic detection of edges for bands of width 8 or higher, whereas the ROA operator detects systematically only from width 13. Some spurious edges are present near the edges in the case of the ROA operator. The experiment indicates that the ROEWA operator has better spatial resolution than the ROA oper-

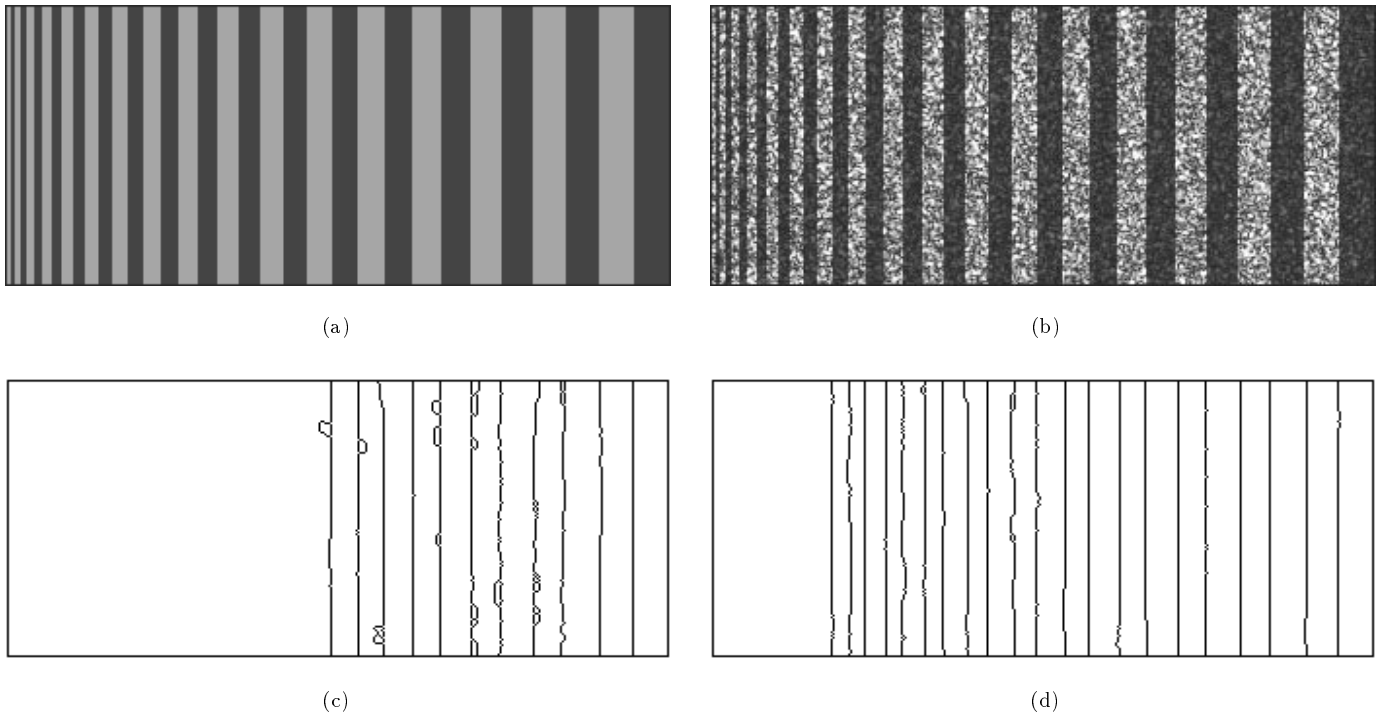


Fig. 5. (a) Ideal image consisting of vertical lines of width 2 to 18 pixels. (b) The simulated single-look speckled image. (c) The segmentation obtained with the ROA edge detector and watershed thresholding. (d) The segmentation obtained with the ROEWA edge detector and watershed thresholding.

ator, for a given speckle reduction capacity. However, we have chosen a very strong smoothing to place ourselves in a multiedge situation. We could of course use a smaller window and detect edges at finer scales with the ROA operator, at the risk of a higher false alarm rate.

Let us now examine a more realistic case. We synthesized the cartoon image shown in amplitude in Fig. 6 (a) by a first order Markov random field with four classes. The reflectivity ratio between subsequent classes is  $6dB$ . This image approximately corresponds to the multiedge model presented in section III-A. The mean region width  $1/\lambda = 13.4$  pixels. Fig. 6 (b) shows the same image multiplied with single-look speckle. The correlation properties of the speckle are the same as in the previous example. To compare the performance of the edge detectors, we use Pratt's figure of merit [18]:

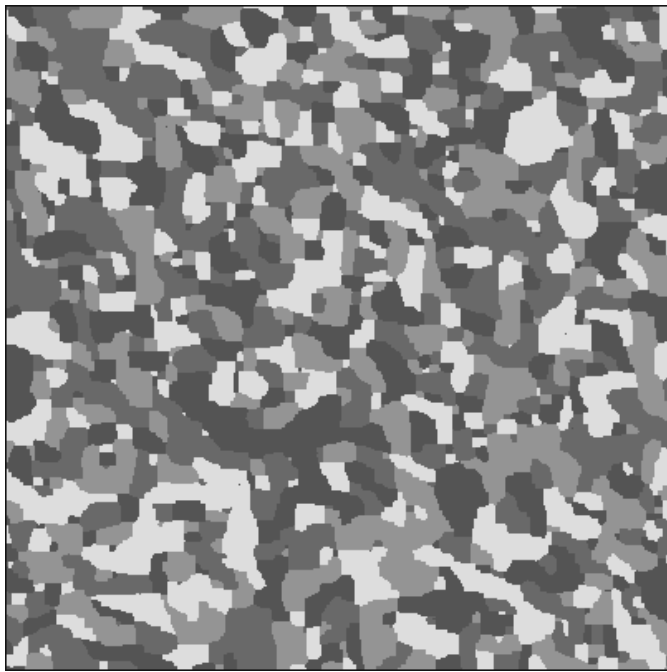
$$\mathcal{P} = \frac{1}{\max\{\mathcal{N}_{DE}, \mathcal{N}_{ID}\}} \sum_{i=1}^{\mathcal{N}_{DE}} \frac{1}{1 + \beta d_i^2},$$

where  $\mathcal{N}_{ID}$  is the number of ideal edge pixels,  $\mathcal{N}_{DE}$  is the number of detected pixels and  $d_i$  is the distance between the  $i$ th detected edge pixel and the closest true edge pixel.  $\beta$  is a calibration constant that is usually set to one. However, as the edges are dense in our test image, so that the nearest ideal edge pixel never is far away, we set  $\beta = 2$  for a stronger penalization of misplaced edge pixels. We accept the closest pixel on each side of a transition as an ideal edge pixel, i.e.,  $d_i = 0$  for every pixel having at least one pixel belonging to another region in its 4-neighborhood. The distance  $d_i$  to an ideal edge for the remaining pixels is obtained as follows:  $d_i = 1$  is attributed to all remaining pixels having one or more pixels with  $d_i = 0$  in their 4-

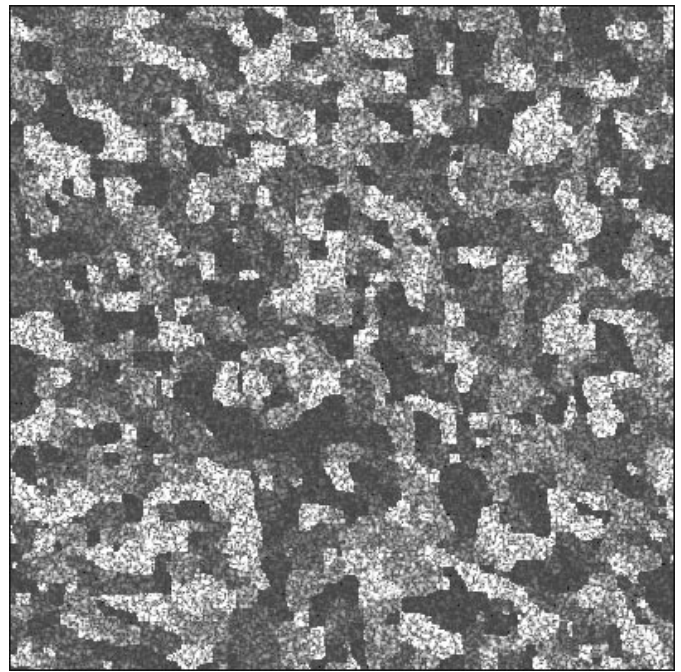
neighborhood. Among the pixels not yet attributed,  $d_i = 2$  is set for every pixel having at least one pixel with  $d_i = 1$  in its 4-neighborhood, and so forth.

Edge strength maps were computed by the ROA operator with window sizes from  $3 \times 3$  to  $19 \times 19$  and by the ROEWA operator with the parameter  $b$  varying over the range 0.1 to 0.8. For each edge strength map the detection threshold maximizing Pratt's figure of merit was determined. Fig. 7 shows the result. The unit along the horizontal axis is the equivalent number of independent pixels in each half of the analyzing window in terms of the speckle reduction obtained by smoothing (see the Appendix). This allows us to compare the results obtained with the ROA operator with different window sizes, with those obtained by the ROEWA operator using exponential weighting functions of varying slope. From Fig. 7 we see that the ROEWA operator yields a better score than the ROA operator over most of the parameter range. However, the difference is relatively small near the maximum of the graphs, and for one window size ( $7 \times 7$ ) the ROA operator performs even better than the ROEWA operator. The difference in favor of the ROEWA operator increases with stronger smoothing. This reflects the fact that the multiedge model is more relevant the larger the analyzing window is. The ROA operator is optimal in the monoedge case, which is more frequently encountered when using small windows. The localization of the maxima of the graphs should not be taken too literally. Such a weak smoothing generally implies an important number of false edges due to speckle. A stronger smoothing gives more meaningful boundaries. The theoretical optimum for the ROEWA operator, according to (7), is  $b = 0.74$ , which corresponds to about 30 independent pixels in each half window.





(a)



(b)

Fig. 6. (a) Ideal image synthesized by a first order Markov random field. (b) The corresponding single-look speckled image.

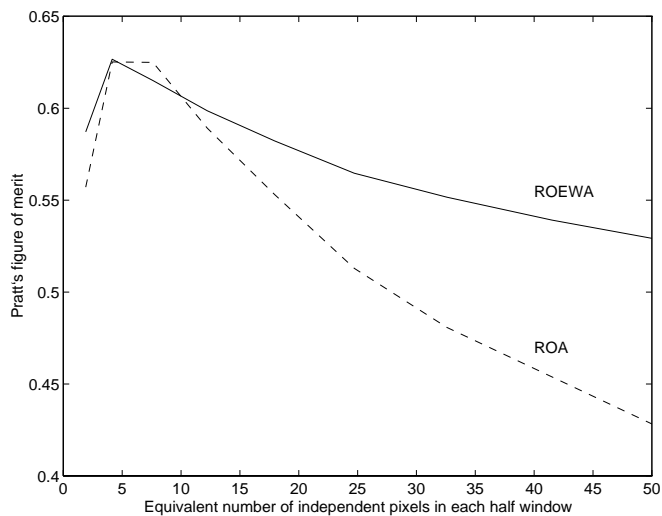


Fig. 7. Pratt's figure of merit for the ROA edge detector with varying window size, and for the ROEWA edge detector with varying slope, applied to the single-look speckled Markov random field image.

Results on real world data are a useful supplement to simulations, but here only a visual appreciation can be given. A multitemporal series of three-look ERS-1 images of an agricultural scene near Bourges, France, was used to test edge detectors and post-processing. An extract of a color composition of 3 dates acquired with monthly intervals is shown in Fig. 8. Note the close resemblance between this scene and the simulated image in Fig. 6. The edge strength maps of the different dates were averaged, supposing that no geometrical changes took place between the acquisitions and that the images are perfectly regis-

tered. Our strategy is to allow a strong over-segmentation in the edge detection step, and then rely on subsequent merging to eliminate false edges. The best results were obtained with a  $13 \times 13$  window for the ROA operator and with  $b = 0.73$  for the ROEWA operator. Given the speckle correlation, the two detectors have about the same speckle reduction capacity with these parameters. Visual inspection of the segmentations revealed only slight differences in favor of the ROEWA operator. We shall use this image to illustrate how complementary post-processing can improve the final result. Fig. 9 shows the initial segmentation, obtained with the ROEWA operator with parameter  $b = 0.73$  and the modified watershed algorithm with threshold 1.53. The threshold was deliberately set very low to make sure that practically all significant edges are detected, resulting in a massively over-segmented image. All the three merging criteria mentioned in section III-E were compared. The LR measure (13) gave the result which agreed best with our conception of the regions. The unequal variance Student's  $t$ -test gave similar results, whereas the classic Student's  $t$ -test performed poorly. In the final segmentation shown in Fig. 10, the number of regions has been reduced from over 5000 to about 600. Adjacent regions for which the log-likelihood  $\ell_{\text{merge}} > -1.85$  for all three dates were merged. The threshold indicates that we accepted some irregularities within the regions. In addition, regions containing only one pixel were supposed to be due to speckle and thus eliminated. The merging order was defined by the iterative pairwise mutually best merge criterion. Almost all regions that we can distinguish by eye have been detected. Some regions still seem to be split in several parts, the edges are sometimes irregular due to speckle, and the corners are slightly rounded due to the strong smoothing used by the edge detector. It is, nevertheless, a surprisingly good SAR image segmentation.

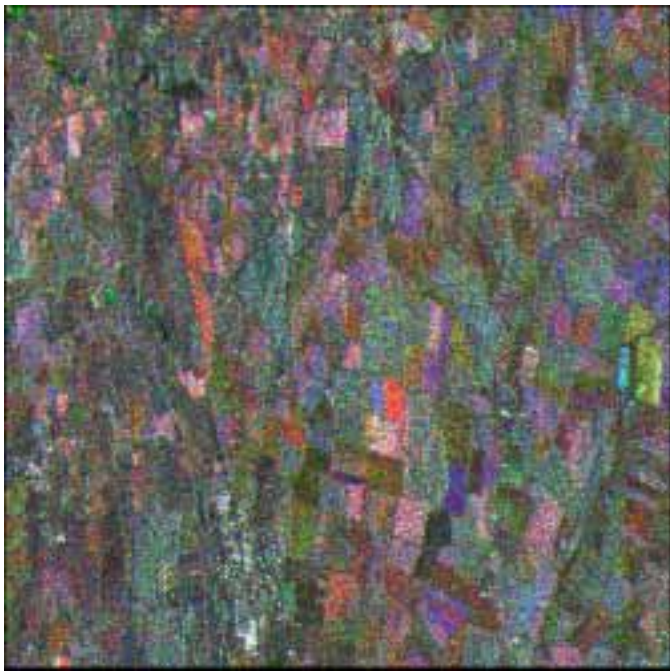


Fig. 8. Extract of a color composition of three SAR images of an agricultural scene near Bourges, France ©ESA - ERS-1 data - 1993, Distribution SPOT Image.

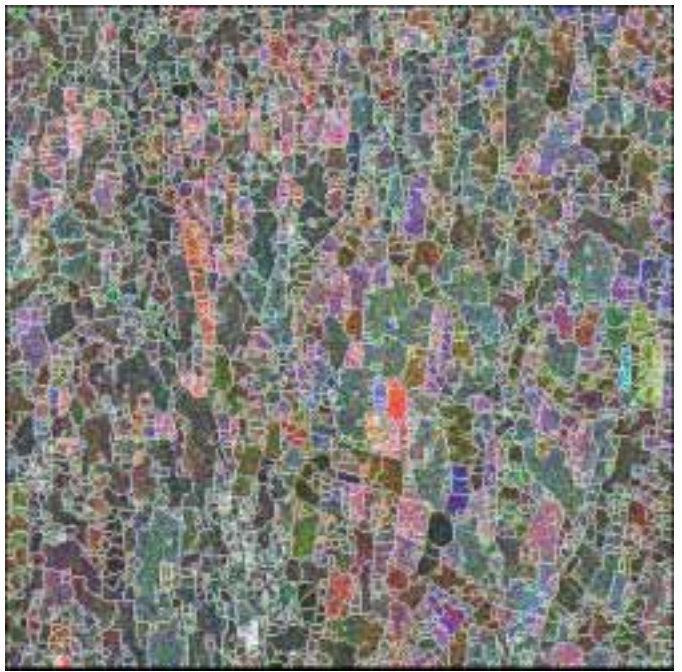


Fig. 9. Over-segmented image obtained by the ROEWA operator and watershed thresholding.

## V. DISCUSSION

The estimator of local means used by the ROEWA operator is optimized for a stochastic multiedge model. We have shown that an exponentially weighted mean with a correctly adjusted slope gives the optimal tradeoff between localization precision and speckle suppression when the reflectivity jumps follow a Poisson distribution. This multiedge model is primarily adapted to describe scenes composed of distinct regions of relatively uniform reflectivity, but of strongly varying size. Exponential weighting is strictly optimal only for scene types which correspond exactly to the stochastic image model. Moreover, we supposed uncorrelated speckle. Equivalent estimators for other scene models and for correlated speckle can be developed by substituting the appropriate spectral density functions into (6), but the impulse response will in general not be any simple, analytic function like the one that we found here.

The arithmetic mean, used by the ROA operator, is the ML estimator of the mean value for a stationary process. The ROA operator is hence spatially optimal in a monoedge context, i.e., when the distance between edges is larger than the width of a half window. If the regions generally are big, as compared to the window size that is necessary to obtain a sufficient speckle suppression, the ROA operator is bound to perform better than the ROEWA operator.

To decide whether or not the ROEWA operator can bring an improvement, as compared to the ROA operator for a given image, several factors must be considered: the average region size and the variations in region size, the contrast between different regions, the ENIL, and the speckle correlation. The ROEWA should theoretically perform better than the ROA operator when the reflectivity approximately corresponds to the multiedge model, the mean re-

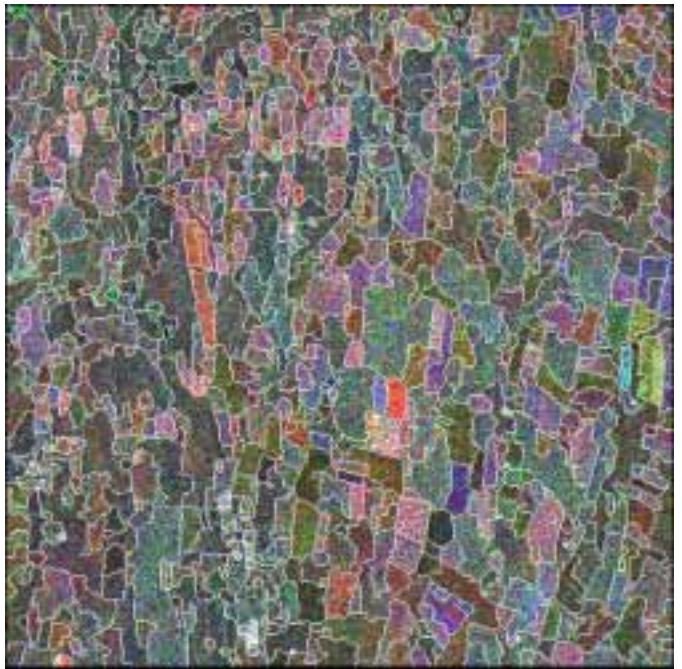


Fig. 10. Segmentation obtained by the ROEWA operator, watershed thresholding and region merging.

gion width is small, and the ENIL is low. With increasing ENIL or mean region width, the monoedge model becomes more appropriate, and the ROA operator can be expected to perform better.

The experimental results confirm the theoretical discussion above. Edge detection on a single-look image composed of vertical bands of gradually increasing width indicate that the ROEWA detector permits a strong speckle reduction without degrading the spatial resolution as much

as the ROA operator. Here we have deliberately placed ourselves in a rather extreme multiedge situation.

On another simulated single-look image, where the reflectivity closely approximates the proposed multiedge model, the ROA and ROEWA detectors were compared over a wide range of window sizes and corresponding slopes of the exponential weighting function, in terms of Pratt's figure of merit. For the smallest windows, the monoedge hypothesis is generally verified and the superiority of the ROA operator is confirmed, even though the scores are very close. With stronger smoothing, corresponding to larger windows, the multiedge model becomes more relevant, and the performance difference in favor of the ROEWA operator increases steadily. Strong smoothing is here necessary to avoid numerous false edges due to speckle, due to the low ENIL and the high speckle correlation.

A hybrid segmentation scheme, which combines the proposed edge detection method with LR region merging, was shown to give excellent results on multitemporal ERS-1 images of an agricultural scene. The difference between the results obtained by the ROA and ROEWA operators was small. This reflects the fact that typical regions are so large that the monoedge model is just as appropriate as the multiedge model for the window size used. Such segmentations can e.g., be used to improve thematic classifications [19]. It should be stressed that this is a very rapid segmentation method. On a Silicon Graphics INDY workstation with a MIPS R4400 200MHz CPU and 64 MB of memory, the ROEWA operator, the watershed thresholding, and the LR region merging needed only 12 s to process three channels of  $512 \times 512$  pixels to produce the result in Fig. 10. This makes our method more than an order of magnitude faster than another sophisticated SAR segmentation scheme, the RWSEG algorithm [5], which is implemented in the CAE-SAR module of the ERDAS IMAGINE software package. The quality of the results are comparable.

## VI. CONCLUSION

In this article, we propose a new CFAR edge detector for SAR images, which is optimal under a stochastic multiedge model. It has been shown to perform better than the ROA operator for images which closely approximate the multi-edge model, especially when the average region width is small and the ENIL is low. The ROEWA operator, watershed thresholding, and LR region merging constitute a very efficient segmentation scheme. The watershed thresholding can be replaced by more advanced edge extraction methods, e.g., based on the powerful concepts of basin dynamics [20] and edge dynamics [21].

The ROEWA operator is a simple, nonadaptive edge detector. There are several other approaches to edge detection and segmentation in a multiedge context. Multiresolution ROA operators [22] combine the ratios computed with different window sizes according to their statistical significance. The ideal solution would be a spatially adaptive LR operator, which varies the window size, the window form and the way it is split, so that the local arithmetic means are estimated on complete, uniform regions. However, these perfectly homogeneous zones are unknown, and difficult to identify in the presence of speckle. The practical solution is to try to iterate towards the best segmentation. The RWSEG algorithm [5], for example, combines edge detection and region growing iteratively. Stochastic

methods based on Markov random fields and simulated annealing [8] iterate towards a segmentation which minimizes a global cost function. Such methods may give even better results, at the cost of a higher computational complexity.

## APPENDIX

Let us suppose the intensity  $I$  to be a wide-sense stationary process. Taking the block-average of  $N$  pixels,  $\hat{\mu}_{Ib} = 1/N \sum_{k=1}^N I_k$ , reduces the variance with a factor  $N$  if the pixels are uncorrelated

$$\sigma_{Ibu}^2 = \frac{\sigma_I^2}{N}$$

If the pixels are correlated

$$\sigma_{Ibc}^2 = \frac{\sigma_I^2}{N^2} \sum_{k=1}^N \sum_{l=1}^N \rho(|k-l|) \quad (14)$$

where  $\rho(m)$ ,  $m \geq 0$ , are the autocorrelation coefficients.

In SAR images, the speckle correlation typically becomes insignificant for distances superior to 2 or 3 pixels. More generally, we may suppose  $\rho(m) = 0$ ,  $m > M$ , and  $M \ll N$ , so that (14) can be rewritten as

$$\sigma_{Ibc}^2 = \frac{\sigma_I^2}{N^2} \left[ N + 2 \sum_{k=1}^M (N-k) \rho(k) \right]. \quad (15)$$

The factor with which the variance is reduced gives us the equivalent number of independent pixels in the analyzing window. Let us now consider the speckle reduction obtained by one half window of the ROEWA operator. We first employ the ISEF  $f$  in one direction

$$\begin{aligned} \sigma_{Ifc}^2 &= \sigma_I^2 \sum_{k=-\infty}^{\infty} \sum_{l=-\infty}^{\infty} f(k)f(l)\rho(k-l) \\ &= \sigma_I^2 \left( \frac{1-b}{1+b} \right)^2 \sum_{k=-\infty}^{\infty} \sum_{m=-M}^M b^{|k|+|k+m|} \rho(m) \\ &= \sigma_I^2 \left( \frac{1-b}{1+b} \right)^2 \sum_{m=-M}^M \left[ |m| + \frac{1+b^2}{1-b^2} \right] b^{|m|} \rho(|m|). \end{aligned}$$

The normalized causal filter  $f_1$  in the perpendicular direction gives

$$\begin{aligned} \sigma_{If_1c}^2 &= \sigma_{Ifc}^2 \sum_{k=0}^{\infty} \sum_{l=0}^{\infty} f_1(k)f_1(l)\rho(k-l) \\ &= \sigma_{Ifc}^2 \frac{(1-b)^2}{1-b^2} \left[ 1 + 2 \sum_{m=1}^M b^{3m} \rho(m) \right]. \end{aligned}$$

The equivalent number of independent pixels in a half window of the ROEWA operator is thus  $\sigma_I^2/\sigma_{If_1c}^2$ , which can be compared to the corresponding number for a half window of the ROA operator obtained by employing (15) in the horizontal and vertical direction.

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